# Selection and Covariance in Galaxy Cluster Surveys: A Multi-λ Model for Local Counts

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# Precise determination of DM virial scaling

virial relation in dark matter from 18 large N-body sims, 6 codes



#### A power-law + scatter model for multiple observables

• For  $i^{\text{th}}$  proxy, mean behavior of  $s_i = \ln(S_i)$  is linear in  $\ln M$  w/ slope  $m_i$ . For N such signals

$$\overline{\mathbf{s}}(\mu, z) = \mathbf{m}(z)\mu + \mathbf{b}(z)$$
  $\mu = \ln M$ 

• assume a log-normal joint likelihood about the mean

$$p(\mathbf{s} \mid \boldsymbol{\mu}, z) = \frac{1}{(2\pi)^{N/2} |\Psi|^{1/2}} \exp[-\frac{1}{2}(\mathbf{s} - \overline{\mathbf{s}})' \Psi^{-1}(\mathbf{s} - \overline{\mathbf{s}})]$$

where  $\Psi$  is the *covariance in signals* at fixed mass and epoch

$$\Psi_{ij} = \left\langle \left( s_i - \overline{s}_i(\mu, z) \right) \left( s_j - \overline{s}_j(\mu, z) \right) \right\rangle$$

Local model for multi-observable counts (the *s*-function )

- locally power-law mass function  $dp = n(\mu) dV$  $n(\mu) = A \exp(-\alpha \mu)$
- convolve with log-normal likelihood for **s** to find the

joint property space density

$$n(\mathbf{s}) = \frac{A\Sigma}{(2\pi)^{(N-1)/2} |\Psi|^{1/2}} \exp\left[-\frac{1}{2} \left(\mathbf{s}' \Psi^{-1} \mathbf{s} - \frac{\overline{\mu}^2(\mathbf{s})}{\Sigma^2}\right)\right]$$

where  $\Sigma^2$  is the mass variance, and  $\mu$  is the log-mean mass

$$\Sigma^2 = \left(\mathbf{m}'\Psi^{-1}\mathbf{m}\right)^{-1}$$
  $\overline{\mu}(\mathbf{s}) =$ 

$$\overline{\mu}(\mathbf{s}) = \frac{\mathbf{m}' \Psi^{-1} \mathbf{s}}{\mathbf{m}' \Psi^{-1} \mathbf{m}} - \alpha \Sigma^2$$

Note: **b**=0 or s = s-b(z) above.

## Explicit 2D number counts: contours of log(n(s))



## Mass selection properties

• Bayes' theorem => Gaussian expectation for mass selection

$$p(\mu \mid \mathbf{s}) = \frac{1}{\sqrt{2\pi}\Sigma} \exp\left[-\frac{\left(\mu - \overline{\mu}(\mathbf{s})\right)^2}{2\Sigma^2}\right]$$

# with *biased mean*

$$\overline{\mu}(\mathbf{s}) = \overline{\mu}_0(\mathbf{s}) - \alpha \Sigma^2$$
selection bias from asymmetric scatter off a steep MF

$$\overline{\mu}_0(\mathbf{s}) = \sum^2 \left( \mathbf{m}' \Psi^{-1} \mathbf{s} \right)$$

inverse of input log-mean relation

<u>Good news</u>: bias in mass scales as the <u>variance</u> <u>Bad news</u>: high-end mass function is <u>steep</u>,  $\alpha \sim 3$ 

# Explicit 2D example

 $\Psi = \begin{bmatrix} \sigma_1^2 & r\sigma_1\sigma_2 \\ r\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}$  correlation coefficient  $r \in (-1,1)$ 

• define equivalent scatter in mass

$$\sigma_{\mu i} = \sigma_i / m_i$$

• the log-mass variance is a harmonic mixture

$$\Sigma^{-2} = \frac{\sigma_{\mu 1}^{-2} + \sigma_{\mu 2}^{-2} - 2r\sigma_{\mu 1}^{-1}\sigma_{\mu 2}^{-1}}{1 - r^2}$$

To minimize scatter and bias in mass, we want:

- small intrinsic scatter
- steep mass relation  $(m_i > 1)$
- anti-correlated signals

 $(\text{if } m_i > 0)$ 





# 17,000 SDSS maxBCG clusters with RASS detections/upper limits

 $\sigma_{\ln L_X, N_{gal}} = 0.83 \pm 0.03$ 

Rykoff et al 2008a



Rykoff et al 2008b

$$\ell = \ln L_{\rm X} ; \sigma_{\mu,\ell} = \sigma_{\ell} / m_{\ell} \qquad \nu = \ln N_{\rm gal} ; \sigma_{\mu,\nu} = \sigma_{\nu} / m_{\nu}$$

• log-mean behavior of binned data with mass

$$\overline{\ell}(\mathbf{v}) = m_{\ell} \Big( \overline{\mu}(\mathbf{v}) + \alpha(\mu) r \sigma_{\mu,\ell} \sigma_{\mu,\nu} \Big)$$

• implied slope of scaling with mean mass may be biased

$$d\overline{\ell}/d\overline{\mu} = m_{\ell} + (r\sigma_{\mu,\ell}\sigma_{\mu,\nu})d\alpha/d\mu$$

• variance is sensitive only to signal correlation

$$\sigma_{\ell,\nu}^{2} = m_{\ell}^{2} \left( \sigma_{\mu,\ell}^{2} + \sigma_{\mu,\nu}^{2} - 2r\sigma_{\mu,\ell}\sigma_{\mu,\nu} \right)$$

# L<sub>X</sub>–M from maxBCG sample

Johston et al 2007 Rykoff et al 2008b

 $M_{200}$  from weak lensing,  $L_{\rm X}$  from RASS, in fixed  $N_{\rm gal}$  bins



Good agreement between X-ray and optically selected samples Non-zero optical-Xray correlation can *tilt*  $N_{\rm gal}$ -binned relation due to running of MF slope  $\alpha(M).$ magnitude scales with  $L_X$ - $N_{gal}$ covariance

# Is a power-law + multivariate Gaussian generic?

#### with Lorena Gazzola, F. Pearce (Nottingham)

Millennium Simulation: Gadget2 with gas under two physical treatments:

– preheating

gravity only

Covariance in ~3000 halos at z=0 with  $M_{200} > 3x10^{13}$  Msun/h

Stanek et al, in prep



#### Local L-T relation: low-hanging covariant fruit?

Nord et al 2008



slopes.

# Selection & Characterization: How to combine approaches?

Method	slope / scatter	mass scatter	blended fraction	comment
Optical	1.0 ± 0.2 / 0.1–0.5 ?	0.1–0.5 ?	5–20%	f <sub>blend</sub> is likely to be z-dependent
SZ	1.6 ± 0.2 / 0.1–0.2 ?	0.06-0.12 ?	5-20% ?	ditto above, + no published detections
X-ray	$1.6 \pm 0.1 / 0.6 \pm 0.1$	$0.37 \pm 0.05$	< 5%	Stanek etal scatter may be high

• SZ + optical will be done jointly (SPT + DES)

Use X-rays to characterize these detections?

- + mostly source photons
- + more clusters with well-measured  $T_X$  (compared to blind)
- timing: SZ source lists not yet available

#### Summary

 cluster survey analysis requires understanding of mass proxies *basic halo model*: power-law mean + log-normal covariance p(slµ) fixed s selects log-normal M dist'n with mean biased by αΣ<sup>2</sup> (co-)variance needs to be understood

\* Apparent variation in the mass scale will bias best-fit cosmology.
\* Variance in the proxy-mass relation will bias mass selection.

- value of multiple cluster measures improved mass selection, understand covariance (physics)
- role of simulations

test robustness of PL+log-normal covariance model selection function from mock survey skies (line-of-sight blending)

• role of XMM?

discovery or characterization? mix of both!